

Statistical Sciences

DoSS Summer Bootcamp Probability Module 2

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Recap

Learnt in last module:

- Measurable spaces
 - Sample Space
 - $\triangleright \sigma$ -algebra
- Probability measures
 - $\,\triangleright\,$ Measures on $\sigma\text{-field}$
 - ▷ Basic results
- Conditional probability
 - ▷ Bayes' rule
 - $\,\triangleright\,$ Law of total probability



Outline

- Independence of events
 - ▷ Pairwise independence, mutual independence
 - ▷ Conditional independence
- Random variables
- Distribution functions
- Density functions and mass functions
- Independence of random variables



Recall the Bayes rule:

$$P(A \mid B) = rac{P(A \cap B)}{P(B)}, \quad P(B) > 0$$

- What if B does not change our belief about A?
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- Equivalently, $P(A \cap B) = P(A)P(B)$.

use this as the defaiter of independence



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- Equivalently, $P(A \cap B) = P(A)P(B)$.

Independence of two events

Two events A and B are independent if $P(A \cap B) = P(A)P(B)$.

Remark:



Consider more than 2 events:

Pairwise independence

We say that events A_1, A_2, \cdots, A_n are pairwise independent if

$$P(A_i \cap A_j) = P(A_i) \cdot P(A_j), \quad \forall i \neq j$$



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Mutual independence

We say that events A_1, A_2, \dots, A_n are mutually independent or independent if for all subsets $l \in \{1, 2, \dots, n\}$

$$P(\cap_{i\in I}A_i)=\prod_{i\in I}P(A_i)$$

$$\begin{array}{c} \mu_{utrul} \circ u \, dypule \implies pairwise independent.\\ chick 1 = \left(\overline{i}, \overline{j}\right), \overline{i} \in \overline{i}^{*}, \\ DRONTO \\ Thum \left(P\left(A_{\overline{i}} \land A_{\overline{i}}\right) = \left(P\left(A_{\overline{c}}\right) \left|P\left(A_{\overline{i}}\right) = u + \frac{1}{2}\right| = \frac{1}{2} \quad 0 < 0 \\ July 10, 2024 \\ 5/20 \end{array}\right)$$

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Remark:



- Toss two fair coins.;
- *A* = { First toss is head}, *B* = { Second toss is head }, *C* = { Outcomes are the same };

•
$$A = \{HH, HT\}, B = \{HH, TH\}, C = \{HH, TT\};$$



- Toss two fair coins.;
- $A = \{$ First toss is head $\}, B = \{$ Second toss is head $\}, C = \{$ Outcomes are the same $\};$ $P(A) = P(B) = P(B) = P(C) = \frac{1}{4} = \frac{1}{2}$
- $A = \{HH, HT\}, B = \{HH, TH\}, C = \{HH, TT\};$
- $P(A \cap B) = P(A)P(B), P(A \cap C) = P(A)P(C), P(B \cap C) = P(B)P(C);$
 - $AOB = \{HM\} \\BNC = \{HM\} \\CAA = \{HH\} \\AB, C are pairwire indeput.$



- Toss two fair coins.;
- *A* = { First toss is head}, *B* = { Second toss is head }, *C* = { Outcomes are the same };
- $A = \{HH, HT\}, B = \{HH, TH\}, C = \{HH, TT\};$
- $P(A \cap B) = P(A)P(B), P(A \cap C) = P(A)P(C), P(B \cap C) = P(B)P(C);$
- $P(A \cap B \cap C) \neq P(A)P(B)P(C)$. $A \cap B \cap C = \{Hu\} \implies P(A \cap B \cap C) = \frac{1}{4} \neq (\frac{1}{2})^3 = P(A)P(B)P(C)$



Conditional independence

Two events A and B are conditionally independent given an event C if

 $P(A \cap B \mid C) = P(A \mid C)P(B \mid C).$



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Conditional independence

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Example:

Previous example continued:

- $A = \{HH, HT\}, B = \{HH, TH\}, C = \{HH, TT\};$
- $P(A \cap B \mid C) = ?, P(A \mid C)P(B \mid C) = ?$
 - $\begin{array}{l} A \ OB = \{ \mu H \}, \\ A \ OB \cap C = \{ \mu H \}, \\ P(A \ OB | c) = \frac{P(A \ OB \ OC)}{P(c)} = \frac{4}{3} : \frac{1}{2}, \\ P(C) = \frac{P(A \ OB \ OC)}{3} = \frac{4}{3} : \frac{1}{2}, \\ P(A \ OC) = \frac{P(A \ OC)}{P(C)} = \frac{4}{3} : \frac{1}{2}, \\ P(B \ OC) = \frac{P(B \ OC)}{P(C)} = \frac{1}{3} : \frac{1}{3}, \\ P(B \ OC) = \frac{P(B \ OC)}{P(C)} = \frac{1}{3} : \frac{1}{3}, \\ P(B \ OC) = \frac{P(B \ OC)}{P(C)} = \frac{1}{3} : \frac{1}{3$



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not conditionably

Conditional independence

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Example:

Previous example continued:

- $A = \{HH, HT\}, B = \{HH, TH\}, C = \{HH, TT\};$
- $P(A \cap B \mid C) = ?, P(A \mid C)P(B \mid C) = ?$

Remark:

Equivalent definition:

$$P(A \mid B, C) = P(A \mid C).$$
 (an show this easily by objection.



Random variables

Idea:

Instead of focusing on each events themselves, sometimes we care more about functions of the outcomes.



Random variables

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Instead of focusing on each events themselves, sometimes we care more about functions of the outcomes.

- Toss a fair coin twice: {*HH*, *HT*, *TH*, *TT*}
- Care about the number of heads: $\{2,1,0\}$



Random variables

Idea:

Instead of focusing on each events themselves, sometimes we care more about functions of the outcomes.

Example:

- Toss a fair coin twice: {*HH*, *HT*, *TH*, *TT*}
- Care about the number of heads: $\{2,1,0\}$



Figure: Mapping from the sample space to the numbers of heads



Random Variables

Example:

- Select twice from red and black ball with replacement: {RR, RB, BR, BB} = \mathcal{G}_{-} .
- Care about the number of red balls: $\{2,1,0\}$



Figure: Mapping from the sample space to the numbers of red balls



Random Variables

Merits:

- Mapping the complicated events on σ -field to some numbers on real line.
- Simplify different events into the same structure





Random Variables

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Random Variables

Consider sample space Ω and the corresponding σ -field \mathcal{F} , for $X : \Omega \to \mathbb{R}$, if

$$A \in \mathcal{R} \quad (ext{Borel sets on } \mathbb{R}) \Rightarrow X^{-1}(A) \in \mathcal{F},$$

then we call <u>X as a random variable</u>. Here $X^{-1}(A) = \{\omega : X(\omega) \in A\}$. We can also say X is \mathcal{F} -measurable.



Distribution functions $(\Omega, \mathcal{F}, \mathbb{R}) \xrightarrow{\times} (\mathbb{R}, \mathcal{R}, \mathbb{C})$

probability on cusar ?

Probability measure $P(\cdot)$ on \mathcal{F} can induce a measure $\mu(\cdot)$ on \mathcal{R} :

Probability measure on ${\cal R}$

We can define a probability μ on (R, \mathcal{R}) as follows:

$$orall A\in \mathcal{R}, \quad \mu(A):=P(X^{-1}(A))=P(X\in A).$$

Then μ is a probability measure and it is called the distribution of X.

You can chick
(i)
$$\mu(\theta) = 0$$

(ii) If A_{1}, A_{2}, \dots are dissiont, $\mu(\bigcup_{i=1}^{W} A_{i}) = \sum_{i=1}^{W} \mu(A_{i})$
(Nii) $\mu(R) = 1$



Probability measure $P(\cdot)$ on \mathcal{F} can induce a measure $\mu(\cdot)$ on \mathcal{R} :

Probability measure on ${\cal R}$

We can define a probability μ on (R, \mathcal{R}) as follows:

$$\forall A \in \mathcal{R}, \quad \mu(A) := P(X^{-1}(A)) = P(X \in A).$$

Then μ is a probability measure and it is called the distribution of X.

Remark:

Verify that μ is a probability measure.

- $\mu(\mathbb{R}) = 1.$
- If $A_1, A_2, \dots \in \mathcal{R}$ are disjoint, then $\mu(\cup_{i=1}^{\infty} A_i) = \sum_{i=1}^{\infty} \mu(A_i)$.



Consider the special set that belongs to \mathcal{R} , $(-\infty, x]$:

Cumulative Distribution Function

The cumulative distribution function of random variable X is defined as follows:

$${\it F}(x):={\it P}(X\leq x)={\it P}(X^{-1}((-\infty,x])),\quad orall x\in \mathbb{R}.$$

$$\left\{ k \notin x \right\} = \left\{ x \in (-\omega, x] \right\}$$
$$= \left\{ x^{-1} \left((-\omega, x] \right) \right\}$$



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Proofs of properties of CDF (first 2 properties):

$$F(x) = IP(x \le x)$$
When $x \to \infty$, $\{x \le x\}$, D .
By the cationity from below,
 $(Im IP(x \le x) = IP(D) = I$.
When $x \to -\infty$, $\{x \le x\}$, D .
By the ontimely from about,
 $Im IP(x \le x) = IP(\varphi) = 0$.
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$$(\operatorname{Proof}: F(x) \text{ 15 nen-decreasy})$$

$$(\operatorname{Lt} \quad x_1 \leq x_2 \dots$$

$$\operatorname{Thm} \quad F(x_1) = \operatorname{IP}(X \leq x_1)$$

$$= (x_2) = \operatorname{IP}(X \leq x_2)$$

$$\operatorname{Sina} \quad x_1 \leq x_2, \{X \leq x_1\} \subset \{X \leq x_2\}$$

$$\operatorname{Thene fore, hy monotonisty of (P, we canche
$$\operatorname{IP}(X \in x_1) \leq \operatorname{IP}(X \leq x_2)$$$$



Classification of the random variables:

- Discrete random variable: X takes either a finite or countable number of possible numbers.
- Continuous random variable: The CDF is continuous everywhere.



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Another perspective (function):

- Discrete random variable: focus on the probability assigned on each possible values
- Continuous random variable: consider the derivative of the CDF (The continuous monotone CDF is differentiable almost everywhere)



Probability mass function

The probability mass function of X at some possible value x is defined by

$$p_X(x) = P(X = x).$$

Relationship between PMF and CDF: If Kis discrete.

$$F(x) = P(X \le x) = \sum_{y \le x} p_X(y)$$
sum of mass function



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Relationship between PMF and CDF:

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Example:

Toss a coin



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Probability density function

The probability density function of X at some possible value x is defined by

$$F(x) = P(X \le x) = \int_{y \le x} f_X(y) \, dy = \int_{-\infty}^x f_X(y) \, dy$$

 $f_X(x) = \frac{d}{dx}F(x)$. Contrinuous case.



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Probability density function

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Relationship between PDF and CDF:

$$F(x) = P(X \le x) = \int_{y \le x} f_X(y) \, dy = \int_{-\infty}^{x} f_X(y) \, dy$$

Example:
$$\chi \sim \mathcal{N}(\mu, \theta^2), \quad f(\chi) = \frac{1}{\sqrt{2\pi} - \theta} \quad \text{erg}\left(- \frac{(\chi - \mu)^2}{2 - \theta^2} \right),$$



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Define independence of random variables based on independence of events:

Independence of random variables

Suppose X_1, X_2, \cdots, X_n are random variables on (Ω, \mathcal{F}, P) , then

 $X_{1}, X_{2}, \dots, X_{n} \text{ are independent}$ $\Leftrightarrow \{X_{1} \in A_{1}\}, \{X_{2} \in A_{2}\}, \dots, \{X_{n} \in A_{n}\} \text{ are independent}, \qquad \forall$ $\Leftrightarrow P(\bigcap_{i=1}^{n} \{X_{i} \in A_{i}\}) = \prod_{i=1}^{n} P(\{X_{i} \in A_{i}\}) \qquad \forall A_{i} \in \mathcal{P},$

$$A_i \in \mathcal{R}$$

Example:

Toss a fair coin twice, denote the number of heads of the *i*-th toss as X_i , then X_1 and X_2 are independent.

- *A_i* can be {0} or {1}
- $\{(0,0),(0,1),(1,0),(1,1)\}$
- $P(\{X_1 \in A_1\} \cap \{X_2 \in A_2\}) = \frac{1}{4}$ $P(\{X_1 \in A_1\} \cap \{X_2 \in A_2\}) = \frac{1}{4}$

•
$$P(\{X_1 \in A_1\}) = P(\{X_2 \in A_2\}) = \frac{1}{2}$$

$$\mathbb{P}(\{k_{i}\in A, \mathcal{F} \cap \{k_{1}\in A_{2}\})$$

$$= \mathbb{P}(x\in A_{1}) \mathbb{P}(x\in A_{2})$$

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Example:

Toss a fair coin twice, denote the number of heads of the *i*-th toss as X_i , then X_1 and X_2 are independent.

- A_i can be {0} or {1}
- $\{(0,0),(0,1),(1,0),(1,1)\}$
- $P({X_1 \in A_1} \cap {X_2 \in A_2}) = \frac{1}{4}$

•
$$P(\{X_1 \in A_1\}) = P(\{X_2 \in A_2\}) = \frac{1}{2}$$

How to check independence in practice?

$$\begin{array}{ccc} (i) & \pi_{i} & \angle O \\ (ii) & O \leq \pi_{i} < 1 \\ (iii') & (\leq \pi_{i}) \end{array}$$



Corollary of independence

If X_1, \cdots, X_n are random variables, then X_1, X_2, \cdots, X_n are independent if

$$P(X_1 \leq x_1, \cdots, X_n \leq x_n) = \prod_{i=1}^n P(X_i \leq x_i)$$



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$$P(X_1 \leq x_1, \cdots, X_n \leq x_n) = \prod_{i=1}^n P(X_i \leq x_i)$$

Remark:

Independence of discrete random variables

Suppose X_1, \dots, X_n can only take values from $\{a_1, \dots\}$, then X_i 's are independent if

$$P(\cap \{X_i = a_i\}) = \prod_{i=1}^n P(X_i = a_i)$$



Problem Set

Problem 1: Give an example where the events are pairwise independent but not mutually independent.

Problem 2: Verify that the measure $\mu(\cdot)$ induced by $P(\cdot)$ is a probability measure on \mathcal{R} .

Problem 3: Prove properties 3 - 5 of CDF $F(\cdot)$.

Problem 4: Bob and Alice are playing a game. They alternatively keep tossing a fair coin and the first one to get a H wins. Does the person who plays first have a better chance at winning?

