



UNIVERSITY OF
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Statistical Sciences

DoSS Summer Bootcamp Probability Module 5

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Recap

Learnt in last module:

- Joint and marginal distributions
 - ▷ Joint cumulative distribution function
 - ▷ Independence of continuous random variables
- Functions of random variables
 - ▷ Convolutions
 - ▷ Change of variables
 - ▷ Order statistics

Outline

- Moments
 - ▷ Expectation, Raw moments, central moments
 - ▷ Moment-generating functions
- Change-of-variables using MGF
 - ▷ Gamma distribution
 - ▷ Chi square distribution
- Conditional expectation
 - ▷ Law of total expectation
 - ▷ Law of total variance

Moments

Intuition: How do the random variables behave on average?

Moments

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Expectation

Consider a random vector X and function $g(\cdot)$, the expectation of $g(X)$ is defined by $\mathbb{E}(g(X))$, where

- Discrete random vector

$$\mathbb{E}(g(X)) = \sum_x g(x) p_X(x),$$

- Continuous random vector in \mathbb{R}^n

$$\mathbb{E}(g(X)) = \int_{\mathbb{R}^n} g(x) dF(x) = \int_{\mathbb{R}^n} f_X(x) dx.$$

Moments

Examples (random variable)

- $X \sim \text{Bernoulli}(p)$: $\mathbb{E}(X) = p \cdot 1 + (1 - p) \cdot 0 = p$.
- $X \sim \mathcal{N}(0, 1)$:

$$\mathbb{E}(X) = \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx = 0.$$

Moments

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Examples (random vector)

- $X_i \sim \text{Bernoulli}(p_i)$, $i = 1, 2$:

$$\mathbb{E} \left((X_1, X_2)^\top \right) = \left((\mathbb{E}(X_1), \mathbb{E}(X_2))^\top \right) = (p_1, p_2)^\top.$$

Moments

Properties:

- $\mathbb{E}(X + Y) = \mathbb{E}(X) + \mathbb{E}(Y)$;
- $\mathbb{E}(aX + b) = a\mathbb{E}(X) + b$;
- $\mathbb{E}(XY) = \mathbb{E}(X)\mathbb{E}(Y)$, when X, Y are independent.

Proof of the first property:

Moments

Raw moments

Consider a random variable X , the k -th (raw) moment of X is defined by $\mathbb{E}(X^k)$, where

- Discrete random variable

$$\mathbb{E}(X^k) = \sum_x x^k p_X(x),$$

- Continuous random variable

$$\mathbb{E}(X^k) = \int_{-\infty}^{\infty} x^k dF(x) = \int_{-\infty}^{\infty} x^k f_X(x) dx.$$

Remark:

Moments

Central moments

Consider a random variable X , the k -th central moment of X is defined by $\mathbb{E}((X - \mathbb{E}(X))^k)$.

Remark:

- The first central moment is 0
- Variance is defined as the second central moment.

Variance

The variance of a random variable X is defined as

$$\text{Var}(X) = \mathbb{E}((X - \mathbb{E}(X))^2) = \mathbb{E}(X^2) - (\mathbb{E}(X))^2.$$

Moments

Another look at the moments:

Moment generating function (1-dimensional)

For a random variable X , the moment generating function (MGF) is defined as

$$M_X(t) = \mathbb{E} \left[e^{tX} \right] = 1 + t\mathbb{E}(X) + \frac{t^2\mathbb{E}(X^2)}{2!} + \frac{t^3\mathbb{E}(X^3)}{3!} + \cdots + \frac{t^n\mathbb{E}(X^n)}{n!} + \cdots$$

Moments

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Compute moments based on MGF:

Moments from MGF

$$\mathbb{E}(X^k) = \frac{d^k}{dt^k} M_X(t) \Big|_{t=0}.$$

Moments

Relationship between MGF and probability distribution:
MGF uniquely defines the distribution of a random variable.

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Example:

- $X \sim \text{Bernoulli}(p)$

$$M_X(t) = \mathbb{E}(e^{tX}) = e^0 \cdot (1 - p) + e^t \cdot p = pe^t + 1 - p.$$

- Conversely, if we know that

$$M_Y(t) = \frac{1}{3}e^t + \frac{2}{3},$$

it shows $Y \sim \text{Bernoulli}(p = \frac{1}{3})$.

Change-of-variables using MGF

Intuition: To get the distribution of a transformed random variable, it suffices to find its MGF first.

Properties:

- $Y = aX + b$, $M_Y(t) = \mathbb{E}(e^{t(aX+b)}) = e^{tb} M_X(at)$.
- X_1, \dots, X_n independent, $Y = \sum_{i=1}^n X_i$, then $M_Y(t) = \prod_{i=1}^n M_{X_i}(t)$.

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Remark:

MGF is a useful tool to find the distribution of some transformed random variables, especially when

- The original random variable follows some special distribution, so that we already know / can compute the MGF.
- The transformation on the original variables is linear, say $\sum_i a_i X_i$.

Change-of-variables using MGF

Example: Gamma distribution

$$X \sim \Gamma(\alpha, \beta),$$

$$f(x; \alpha, \beta) = \frac{x^{\alpha-1} e^{-\beta x} \beta^\alpha}{\Gamma(\alpha)} \quad \text{for } x > 0 \quad \alpha, \beta > 0.$$

Compute the MGF of $X \sim \Gamma(\alpha, \beta)$ (details omitted),

$$M_X(t) = \left(1 - \frac{t}{\beta}\right)^{-\alpha} \quad \text{for } t < \beta, \text{ does not exist for } t \geq \beta.$$

Change-of-variables using MGF

Example: Gamma distribution

Observation:

The two parameters α, β play different roles in variable transformation.

- Summation:

If $X_i \sim \Gamma(\alpha_i, \beta)$, and X_i 's are independent, then $T = \sum_i X_i \sim \Gamma(\sum_i \alpha_i, \beta)$.

If $X_i \sim \text{Exp}(\lambda)$ (this is equivalently $\Gamma(\alpha_i = 1, \beta = \lambda)$) distribution, and X_i 's are independent, then $T = \sum_i X_i \sim \Gamma(n, \lambda)$.

- Scaling:

If $X \sim \Gamma(\alpha, \beta)$, then $Y = cX \sim \Gamma(\alpha, \frac{\beta}{c})$.

Change-of-variables using MGF

Example: χ^2 distribution

χ^2 distribution

If $X \sim \mathcal{N}(0, 1)$, then X^2 follows a $\chi^2(1)$ distribution.

Find the distribution of $\chi^2(1)$ distribution

- From PDF: (Module 4, Problem 2)

For X with density function $f_X(x)$, the density function of $Y = X^2$ is

$$f_Y(y) = \frac{1}{2\sqrt{y}}(f_X(-\sqrt{y}) + f_X(\sqrt{y})), \quad y \geq 0,$$

this gives

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} y^{-\frac{1}{2}} \exp\left(-\frac{y}{2}\right).$$

Change-of-variables using MGF

Find the distribution of $\chi^2(1)$ distribution (continued)

- From MGF:

$$\begin{aligned}M_Y(t) &= \mathbb{E}(e^{tX^2}) = \int_{-\infty}^{\infty} \exp(tx^2) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \\&= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2(1-2t)^{-1}}\right) dx \\&= (1-2t)^{-\frac{1}{2}} \int_{-\infty}^{\infty} \mathcal{N}(0, (1-2t)^{-1}) dx, \quad t < \frac{1}{2} \\&= (1-2t)^{-\frac{1}{2}}, \quad t < \frac{1}{2}.\end{aligned}$$

By observation, $\chi^2(1) = \Gamma(\frac{1}{2}, \frac{1}{2})$.

Change-of-variables using MGF

Generalize to the $\chi^2(d)$ distribution

$\chi^2(d)$ distribution

If $X_i, i = 1, \dots, d$ are i.i.d $\mathcal{N}(0, 1)$ random variables, then $\sum_{i=1}^d X_i^2 \sim \chi^2(d)$.

By properties of MGF, $\chi^2(d) = \Gamma(\frac{d}{2}, \frac{1}{2})$, and this gives the PDF of $\chi^2(d)$ distribution

$$\frac{x^{\frac{d}{2}-1} e^{-\frac{x}{2}}}{2^{\frac{d}{2}} \Gamma(\frac{d}{2})} \quad \text{for } x > 0.$$

Conditional expectation

From expectation to conditional expectation:

How will the expectation change after conditioning on some information?

Conditional expectation

From expectation to conditional expectation:

How will the expectation change after conditioning on some information?

Conditional expectation

If X and Y are both discrete random vectors, then for function $g(\cdot)$,

- Discrete:

$$\mathbb{E}(g(X) \mid Y = y) = \sum_x g(x) p_{X|Y=y}(x) = \sum_x g(x) \frac{P(X = x, Y = y)}{P(Y = y)}$$

- Continuous:

$$\mathbb{E}(g(X) \mid Y = y) = \int_{-\infty}^{\infty} g(x) f_{X|Y}(x|y) dx = \frac{1}{f_Y(y)} \int_{-\infty}^{\infty} g(x) f_{X,Y}(x, y) dx.$$

Conditional expectation

Properties:

- If X and Y are independent, then

$$\mathbb{E}(X \mid Y = y) = \mathbb{E}(X).$$

- If X is a function of Y , denote $X = g(Y)$, then

$$\mathbb{E}(X \mid Y = y) = g(y).$$

Sketch of proof:

Conditional expectation

Remark:

By changing the value of $Y = y$, $\mathbb{E}(X \mid Y = y)$ also changes, and $\mathbb{E}(X \mid Y)$ is a random variable (the randomness comes from Y).

Conditional expectation

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Total expectation and conditional expectation

Law of total expectation

$$\mathbb{E}(\mathbb{E}(X \mid Y)) = \mathbb{E}(X)$$

Proof: (discrete case)

Conditional expectation

Total variance and conditional variance

Conditional variance

$$\text{Var}(Y | X) = \mathbb{E}(Y^2 | X) - (\mathbb{E}(Y | X))^2.$$

Conditional expectation

Total variance and conditional variance

Conditional variance

$$\text{Var}(Y | X) = \mathbb{E}(Y^2 | X) - (\mathbb{E}(Y | X))^2.$$

Law of total variance

$$\text{Var}(Y) = \mathbb{E}[\text{Var}(Y | X)] + \text{Var}(\mathbb{E}[Y | X]).$$

Remark:

Problem Set

Problem 1: Prove that $\mathbb{E}(XY) = \mathbb{E}(X)\mathbb{E}(Y)$ when X and Y are independent.

(Hint: simply consider the continuous case, use the independent property of the joint pdf)

Problem 2: For $X \sim \text{Uniform}(a, b)$, compute $\mathbb{E}(X)$ and $\text{Var}(X)$.

Problem 3: Determine the MGF of $X \sim \mathcal{N}(\mu, \sigma^2)$.

(Hint: Start by considering the MGF of $Z \sim \mathcal{N}(0, 1)$, and then use the transformation $X = \mu + \sigma Z$)

Problem Set

Problem 4: The citizens of Remuera withdraw money from a cash machine according to $X = 50, 100, 200$ with probability $0.3, 0.5, 0.2$, respectively. The number of customers per day has the distribution $N \sim \text{Poisson}(\lambda = 10)$. Let $T_N = X_1 + X_2 + \cdots + X_N$ be the total amount of money withdrawn in a day, where each X_i has the probability above, and X_i 's are independent of each other and of N .

- Find $\mathbb{E}(T_N)$,
- Find $\text{Var}(T_N)$.